

2 **Familiarity and preference for pitch probability profiles**

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6 **Abstract** We investigated familiarity and preference
7 judgments of participants toward a novel musical system.
8 We exposed participants to tone sequences generated from
9 a novel pitch probability profile. Afterward, we either
10 asked participants to identify more familiar or we asked
11 participants to identify preferred tone sequences in a two-
12 alternative forced-choice task. The task paired a tone se-
13 quence generated from the pitch probability profile they
14 had been exposed to and a tone sequence generated from
15 another pitch probability profile at three levels of distinc-
16 tiveness. We found that participants identified tone se-
17 quences as more familiar if they were generated from the
18 same pitch probability profile which they had been exposed
19 to. However, participants did not prefer these tone se-
20 quences. We interpret this relationship between familiarity
21 and preference to be consistent with an inverted U-shaped
22 relationship between knowledge and affect. The fact that
23 participants identified tone sequences as even more famil-
24 iar if they were generated from the more distinctive
25 (caricatured) version of the pitch probability profile which
26 they had been exposed to suggests that the statistical
27 learning of the pitch probability profile is involved in
28 gaining of musical knowledge.

30 **Keywords** Music cognition · Pitch probability profile ·
31 Statistical learning · Preference · Familiarity

Introduction

Because affect for music is likely related to the violation
and confirmation of expectancy (Meyer 1956), music psy-
chologists have been motivated to investigate the origins of
listeners' expectancies. It has been argued that the ex-
pectancies are based on the knowledge of statistical regu-
larities in music, which we acquire unintentionally
(without intention) by being exposed to the music (Huron
2006; Tillmann et al. 2000). The mere exposure effect
proposed by Zajonc (1968) relates affect for music to ex-
posure directly. In essence, it proposes that exposure alone
leads to "attitude enhancement" (Zajonc 1968). This sug-
gests that affect for music is dependent on the amount of
exposure. If expectancies are based on the knowledge of
music, and if we assume that exposure to music leads to
knowledge about this music, then both Meyer (1956) and
Zajonc (1968) link what we like to what we know. Exposure
could lead to the knowledge of statistical regularities (for
instance, through unintentional learning). The statistical
regularities could be the basis for expectancies formed by
the listener. The tension a listener feels from violation and
confirmation of those expectancies leads to changes in af-
fect. The hypothesized relationships between exposure,
knowledge of statistical regularities, expectancies, and af-
fect are visualized in Fig. 1. The dashed arrow denotes the
relationship proposed by Zajonc (1968). The mere exposure
effect assumes change in affect through change in exposure.
Meyer (1956) and later Huron (2006) proposed that affect is
related to violation and confirmation of expectancies. This
is denoted by the dotted arrow. The first question we were
interested in exploring is whether participants are able to
gain knowledge of statistical regularities through exposure.

Some have argued that the mere exposure effect is a
form of implicit memory, i.e., that what we like is basically

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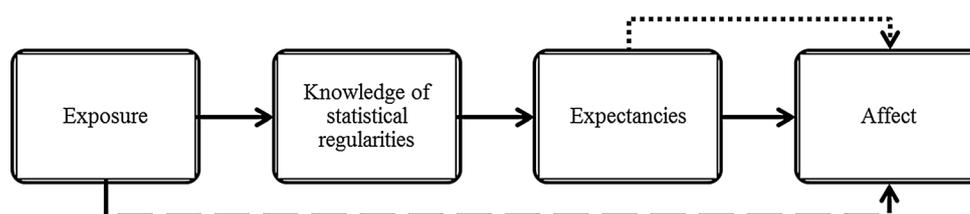
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Fig. 1 Hypothesized relationship between exposure, knowledge of statistical regularities, expectancies, and affect



66 what we know (Seamon et al. 1995; Peretz et al. 1998).
 67 This suggests that asking listeners about their affect for
 68 music assesses their memory implicitly, while asking lis-
 69 teners about their knowledge of music would do so ex-
 70 plicitly. This is the second question we wanted to explore
 71 in our study. However, in a recent study by Loui et al.
 72 (2010), dissociation between knowledge and affect was
 73 reported. In this study, participants “recognized” previ-
 74 ously unheard music fashioned after the same statistical
 75 rules that were also found in the music that they had heard,
 76 but did not display a mere exposure effect. In a study by
 77 Kuhn and Dienes (2005), the opposite pattern was ob-
 78 served. After exposure to a tone sequence set generated
 79 using specific rules, participants liked tone sequences
 80 generated using the same rules more than tone sequences
 81 that violated those rules. However, when asked to differ-
 82 entiate between tone sequences that followed the rules and
 83 tone sequences that violated the rules, participants were
 84 unable to do so. Both these studies provide arguments
 85 against the proposition that the mere exposure effect is a
 86 form of implicit memory as familiarity, and liking ratings
 87 were different. On the other hand, in a study by Schel-
 88 lenberg et al. (2008), it was found that liking as well as
 89 familiarity ratings increase linearly as a function of expo-
 90 sure if participants listen to music unintentionally. Simul-
 91 taneously, recognition ratings for the music increased with
 92 exposure, therefore suggesting a link between mere expo-
 93 sure effect and memory. But liking of music followed an
 94 inverted U-shaped function of exposure if listeners focused
 95 on the music. An inverted U-shaped relationship between
 96 stimulus complexity and liking has been proposed before
 97 by Berlyne (1974). It could be argued that stimulus com-
 98 plexity declines with exposure, as participants become in-
 99 creasingly familiar with stimuli after exposure, such that
 100 overexposure leads to boredom (Cantor 1968). Huron
 101 (2006) proposes that predictability in music, which would
 102 require familiarity with the music obtained through expo-
 103 sure, leads to positive responses. However, high pre-
 104 dictability might also lead to negative appraisal, if the
 105 listener feels as though he or she is being “treated as
 106 musical children” (Huron 2006, p. 141).

107 It should be noted that the form of the relationship be-
 108 tween variables like stimulus complexity or novelty and
 109 preference varies between different stimulus domains

(Martindale et al. 1990). While the study by Schellenberg
 et al. (2008) mentioned above found an inverted U-shaped
 function of exposure to musical pieces and preference, one
 can find positive and negative linear functions of acoustic
 qualities such as loudness or pitch (Martindale et al. 1990),
 and other factors, such as meaningfulness of the stimuli,
 rather than stimuli complexity, may contribute to prefer-
 ence ratings to a greater degree (Martindale and Moore
 1989).

To address the questions of (a) whether participants are
 able to acquire knowledge about statistical regularities of a
 musical system after exposure, and (b) whether results
 differ if we assess participants memory implicitly or ex-
 plicitly, we exposed participants to music from a pre-de-
 fined musical system. To circumvent influence by what
 participants already know, we assessed participants on their
 knowledge of and affect for novel musical systems. The
 statistical regularities of the systems that we used were
 described by Huron and Veltman (2006). Unlike the sta-
 tistical regularities underlying the musical system used by
 Loui et al. (2010), the statistical regularities of the new
 musical system used here were only represented in terms of
 their first-order statistics, that is, as the probability distri-
 bution of pitch occurrences, without accounting for the
 sequence in which they occur. First-order probability pro-
 files have been used in algorithms used to determine the
 musical system of a piece (Krumhansl 1990; Huron and
 Veltman 2006). First-order probability profiles based on
 real music also seem to correspond to the mental hierarchy
 of pitch classes (Krumhansl 1985).

Another variable we sought to modify is the listener’s
 expectancies. By altering the first-order probability profile
 such that it varies in distinctiveness, we hoped to ma-
 nipulate the amount of violation and confirmation of po-
 tentially formed expectancies. We created three levels of
 distinctiveness. The standard level of distinctiveness cor-
 responded to the probability profile described by Huron and
 Veltman (2006). By sharpening the probability profile, we
 created a profile with a high level of distinctiveness. By
 flattening the probability profile, we created a profile with a
 low level of distinctiveness. Note that these manipulations
 kept the rank order of pitch occurrence intact. We exposed
 participants to tone sequences at the standard level of
 distinctiveness. Afterward, we tested either the

154 participants' knowledge of or their affect for tone sequences at all levels of distinctiveness in a two-alternative
 155 forced-choice task. We did this by either asking participants to indicate which melody they found more fam-
 156 ilar (knowledge) or which melody they preferred (affect). Expectancies formed during exposure (i.e., based
 157 on music at the standard level of distinctiveness) are more likely to be confirmed if the participant is tested with tone
 158 sequences at the high level of distinctiveness: Pitch classes occurring often at the standard level, which are the pitch
 159 classes that the participant would expect, occur even more often at the high level of distinctiveness. These expectancies
 160 are more likely to be violated if the participant is tested with tone sequences at the low level of distinctiveness:
 161 Pitch classes occurring often at the standard level, which are the pitch classes that the participant would expect,
 162 occur less at the low level of distinctiveness. The question we addressed with this manipulation is whether distinctiveness
 163 influences participants' responses.

173 By creating two experimental groups that differed in the instructions they received, we aimed to test whether any
 174 knowledge that might be acquired during exposure can be acquired unintentionally, i.e., without the intention to learn.
 175 Participants were focused on the musical stimuli in both conditions. However, while participants in the group with
 176 more instructions were told that they would be asked questions about the music they were about to hear, i.e.,
 177 about the two-alternative forced-choice task, participants in the group with less instructions were not told about that
 178 phase in the experiment. Participants in the group with less instructions thus had no reason to actively try to learn about
 179 the music during the exposure phase.

186 While the mere exposure effect offers no direct prediction, one could argue that participants have not been
 187 exposed to tone sequences at levels of distinctiveness other than the standard level prior to the testing of knowledge or
 188 affect. This would mean lower preference for tone sequences at levels of distinctiveness other than the standard
 189 level. However, if participants acquire knowledge of statistical regularities and generalize the hierarchy of pitch
 190 classes, tone sequences at high levels of distinctiveness could be regarded as more familiar. Researchers who have
 191 proposed that the mere exposure effect is a form of implicit memory would then assume that those tone sequences lead
 to higher preference. A different pattern is expected if there is indeed dissociation between knowledge and affect.

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 194 proposed that the mere exposure effect is a form of implicit memory would then assume that those tone sequences lead
 195 to higher preference. A different pattern is expected if there is indeed dissociation between knowledge and affect.
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200 Methods

201 Stimuli

202 Tone sequences were generated from a version of Temperley's (2007) pitch model using pitch profiles of two
 203 novel musical systems. The two novel musical systems used in this study were two Gregorian chant modes: Hypo-
 204 phrygian and Lydian. The pitch profile of these modes has been described by Huron and Veltman (2006). Huron
 205 and Veltman (2006) tallied the pitch classes occurring in a sample of the Liber Usualis, a book containing over 2000
 206 Gregorian chants. Three versions of pitch profiles were constructed by raising the profile described by Huron and
 207 Veltman (2006) to different exponents using an algorithm by Smith and Schmuckler (2004). An exponent of 1 created
 208 the standard level of distinctiveness. An exponent of 2 created the high level of distinctiveness. An exponent of .5
 209 created the low level of distinctiveness. Figure 2 visualizes the pitch profile for both modes at each level of distinctiveness.
 210 A profile at the high level of distinctiveness is characterized by higher "peaks" and lower "dips." A profile at the low
 211 level of distinctiveness is "flatter" than profiles at higher levels of distinctiveness. Each tone sequence was 15 s long
 212 and consisted of 50 tones of equal duration. Tones were of a flute-like timbre. Tone sequences were generated by
 213 independently drawing from the probability profiles. The range of pitches used for the tone
 214 sequences was C, D, E, F, G, A, B_b, B.
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Fig. 2 Pitch probability profiles for **a** Hypophrygian and for **b** Lydian tone sequences. The low level of distinctiveness is indicated with *dotted lines*. The standard level of distinctiveness is indicated with *dashed lines*. The high level of distinctiveness is indicated with *solid lines*

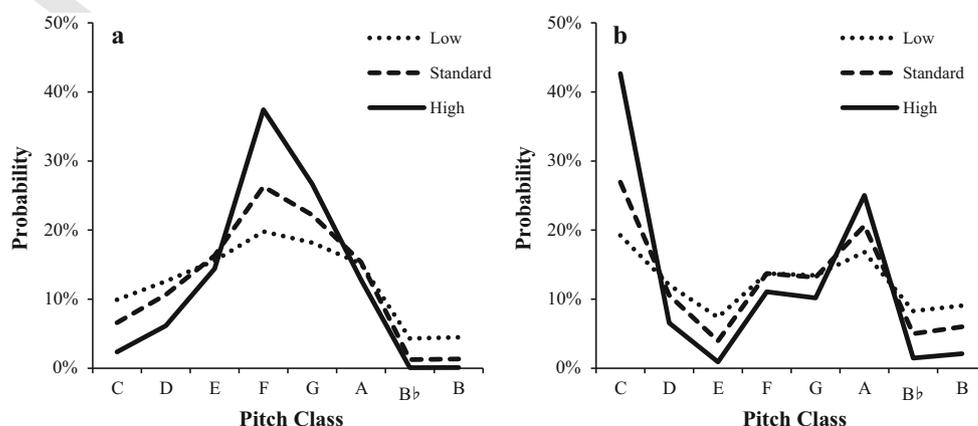
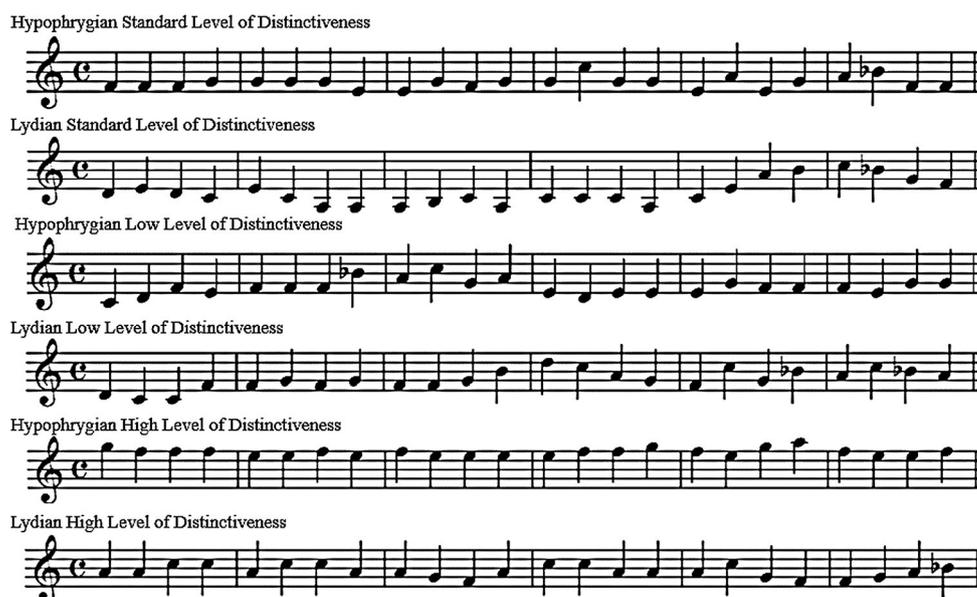


Fig. 3 Excerpts from sample tone sequences



226 sequences was determined by a normal probability density
 227 function with a μ_1 , which was drawn from a normal dis-
 228 tribution with $\mu_2 = 68$ and $\sigma_2 = 3.63$, and $\sigma_1 = 5.39$,
 229 with 68 corresponding to a central pitch of $G\#^4$ and the
 230 interval of $\mu_2 \pm \sigma_1$ spanning the range from $D\#^4$ to $C\#^5$.
 231 An excerpt from example tone sequences for each of the
 232 six probability profiles (Hypophrygian standard, high, and
 233 low level of distinctiveness, Lydian standard, high, and low
 234 level of distinctiveness) can be found in Fig. 3.

235 Participants

236 Eighty-two participants with less than five years of formal
 237 music training were recruited from Queen's University and
 238 compensated monetarily for their time. Twenty-eight of
 239 those participants were male, 54 female.

240 Procedure and Design

241 Testing took place in a sound-attenuated chamber. Par-
 242 ticipants were instructed to adjust the volume to their
 243 preferred level. Testing was divided into two phases. In the
 244 first phase (the exposure phase), participants were exposed
 245 to tone sequences generated from the pitch profile of one
 246 mode (exposed mode) at the standard level of distinctiv-
 247 ness. The mode (either Hypophrygian or Lydian) was
 248 chosen randomly. The exposure phase consisted of 100
 249 individual tone sequences. The exposure phase lasted
 250 25 min. The ensuing second phase (the test phase) com-
 251 prised 30 trials of a two-alternative forced-choice task.
 252 Tone sequences presented during the test phase were new,
 253 i.e., participants had not heard them prior. We paired a tone
 254 sequence from the exposed mode and a tone sequence

generated from the non-exposed mode. There were 10 such
 pairings (trials) at each of the three levels of distinctiv-
 ness; thus, there were a total of 30 trials. The order in
 which the 30 trials were presented was randomized. The
 level of distinctiveness was our within-subject factor.

There were two between-subject factors yielding four
 experimental groups. The first factor was instruction: par-
 ticipants were either informed at the beginning of the ex-
 periment, that the test phase would take place (with
 instruction), or they were told to follow instructions "as
 they come along" (no instruction). The latter group was
 thus unaware of the test phase until it started. The second
 factor divided participants into a familiarity and a prefer-
 ence group. In the test phase, the familiarity group was
 asked to indicate which tone sequence they found more
 familiar. The preference group was asked to indicate which
 tone sequence they preferred. Nineteen participants were
 in the familiarity—no instruction group, 23 participants were
 in the familiarity—with instruction group, 20 participants
 were in the preference—no instruction group, and 20 par-
 ticipants were in the preference—with instruction group.

276 Results

277 The dependent variables were calculated by dividing the
 278 number of times the tone sequence of the exposed mode
 279 was chosen in the test phase by the number of trials at each
 280 level of distinctiveness. The choice of the exposed mode
 281 was classified as "correct." The numbers are expressed as a
 282 percentage of correct responses in Fig. 3. Thus, a percent
 283 "correct" of 50 % would indicate chance performance.
 284 There was no effect of "instruction," $F(1,78) = 1.10$,

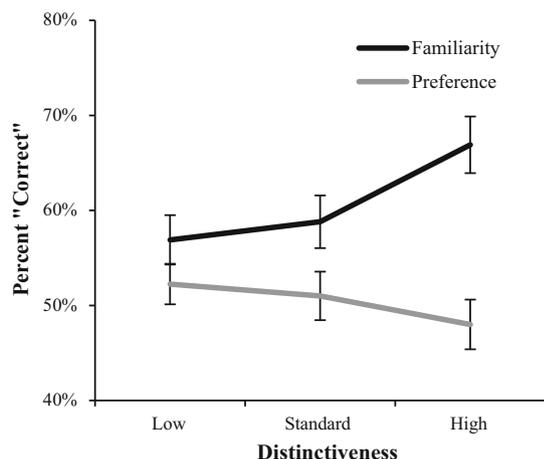


Fig. 4 Percent “correct” for familiarity and preference group at each level of distinctiveness. The choice of the exposed mode was classified as “correct.” Error bars depict the standard error of mean

285 $p = .297$, and no interaction effects involving “instruc-
 286 tion,” $p > .05$, so data were pooled across this factor for
 287 further analysis. There was a main effect of familiarity/
 288 preference, $F(1,80) = 19.24$, $p < .001$, such the percent
 289 “correct” scores of participants in the familiarity group
 290 were higher than the percent “correct” scores of partici-
 291 pants in the preference group. Participants in the fami-
 292 larity group performed higher than chance (one-sample
 293 t -tests, low distinctiveness $t(41) = 2.65$, $p = .011$, stan-
 294 dard distinctiveness $t(41) = 3.18$, $p = .003$, high distinc-
 295 tiveness $t(41) = 5.67$, $p < .001$), but those in the
 296 preference group did not ($ps > .05$). Moreover, there was a
 297 significant interaction between familiarity/preference and
 298 distinctiveness, $F(2,160) = 4.57$, $p = .012$, and no sig-
 299 nificant main effect of distinctiveness, $F(2,160) = 0.81$,
 300 $p = .447$. The linear contrast for distinctiveness in the fami-
 301 larity group was significant ($F(1,41) = 6.11$, $p = .018$),
 302 but not in the preference group ($F(1,39) = 1.70$, $p = .200$).
 303 The data are depicted in Fig. 4.

304 Discussion

305 The participants in the familiarity group indicated that tone
 306 sequences generated from the pitch probability profile of
 307 the mode that they had previously been exposed to were
 308 more familiar. Thus, it can be argued that they generalized
 309 the pitch probability profile they were exposed to: partici-
 310 pants indicated familiarity for tone sequences generated
 311 from the same pitch probability profile as the sequences
 312 heard during exposure. The only similarity between the
 313 tone sequences participants had heard and the tone se-
 314 quences in the test phase is that they were generated from
 315 the same pitch probability profile. As there was no effect of

“instruction,” this demonstrates successful unintentional 316
 learning of a first-order probability profile, i.e., learning 317
 that took place without the intention to learn. Although the 318
 participants in the “with instruction” group were not told 319
 that the tone sequences in the test phase would be similar to 320
 the tone sequences in the exposure phase, it can be argued 321
 that participants may have suspected this. Participants 322
 could then have tried to actively “learn” the pitch prob- 323
 ability profile. The participants in the “no instruction” 324
 group were told to follow instructions “as they come 325
 along.” As they were not told about the test phase, we 326
 assume that any learning that they exhibited took place 327
 unintentionally. Moreover, participants in the familiarity 328
 group chose tone sequences from the pitch probability 329
 profile of the exposed mode at the high level of distinctiv- 330
 eness more often than chance. These tone sequences 331
 followed the statistical rules set forth during exposure more 332
 distinctly. Participants in the familiarity group also chose 333
 tone sequences from the pitch probability profile of the 334
 exposed mode at the low level of distinctiveness. However, 335
 this effect was less pronounced than when tone sequences 336
 were generated at the standard or high level of distinctiv- 337
 eness. The significant linear contrast such that the effect 338
 increased with levels of distinctiveness can be described as 339
 a “caricature effect”: the caricatured version of the prob- 340
 ability profile (the probability profile at the high level of 341
 distinctiveness) appears to be more familiar. This indicates 342
 that participants generalized the probability profile and 343
 supported the theory that salient pitches are important for 344
 acquisition of pitch profiles and the tonal hierarchy 345
 (Krumhansl and Cuddy 2010). In future studies, more 346
 levels of distinctiveness could be tested to validate this idea 347
 of a “caricature effect.” It should be noted that based on 348
 our data, the representation of the distribution is not ab- 349
 solute, so judgments of familiarity are facilitated by the 350
 caricature of the distribution. It is not impossible that 351
 participants abstracted an interval distribution rather than a 352
 pitch distribution. We believe that our argument of dis- 353
 tinctiveness would hold nonetheless, as the distribution of 354
 intervals would become more distinct at the high level of 355
 distinctiveness of the pitch distribution and less distinct at 356
 the low level of distinctiveness of the pitch distribution. 357

The absence of a significant linear contrast of distinc- 358
 tiveness in the preference group, the fact that their behavior 359
 in the test phase was not different from chance, and particu- 360
 larly the strong interaction between distinctiveness and 361
 familiarity/preference do not suggest a positive linear re- 362
 lationship between knowledge and affect. This goes against 363
 the prediction made by mere exposure effect theories: in 364
 our experiment, mere exposure did not lead to “attitude 365
 enhancement” (Zajonc 1968). This also does not evidently 366
 support the idea that affect assessment assesses implicit 367
 memory (Seamon et al. 1995; Peretz et al. 1998). However, 368

369 the findings in our study could be interpreted as assessing
370 implicit memory if the relationship between knowledge
371 and affect follows an inverted U-shaped curve.

372 Szpunar et al. (2004) for instance have suggested that
373 liking follows an inverted U-shaped function of exposure.
374 Participants in our study might have been overexposed to
375 the pitch probability profile, thereby increasing familiarity,
376 but decreasing preference. This decrease was especially
377 visible for tone sequences at the high level of distinctiveness.
378 If exposure is positively correlated with familiarity,
379 then the data could be interpreted as lying on the right part
380 of an inverted U-shaped relationship. Figure 5 graphs the
381 data shown in Fig. 4 in a different fashion: in this figure,
382 the mean percentages in the familiarity group are plotted
383 against the mean percentages in the preference group. As
384 can be seen, the interpretation of the data as lying on the
385 right shoulder of an inverted U-shaped function of exposure
386 could be considered: the more familiar the stimuli
387 were, the less they were preferred. If stimulus complexity
388 declines with familiarity, the interpretation of the data as
389 lying on the left shoulder of an inverted U shape of stimulus
390 complexity could be considered (Berlyne 1974).
391 This would also be consistent with the boredom effect
392 described by Cantor (1968).

393 Another perspective from which to view our results is in
394 reference to schema and prototype theory. Smith and Melara
395 (1990) propose that listeners may understand a musical
396 phrase by treating it as a categorization problem. In our
397 experiment, then the standard probability profiles are the
398 prototypes to which new tone sequences are assimilated.
399 Tone sequences generated from the low level of

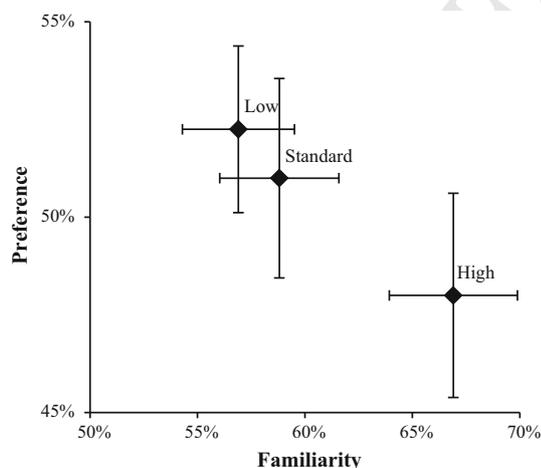


Fig. 5 Percent “correct” at each level of distinctiveness. The percent “correct” of the familiarity group is plotted against the percent “correct” of the preference group for each level of distinctiveness. The horizontal and vertical error bars depict standard error of the mean for the familiarity group and for the preference group, respectively

400 distinctiveness profile require greater effort to be as-
401 similated and thus appear less familiar. In an experiment by
402 Smith and Melara (1990), participants rated musical
403 phrases with varying degrees of prototypicality of Western
404 music. Musical phrases with a greater degree of proto-
405 typicality were rated as less unusual by musically untrained
406 participants. For the same participants, ratings of pleas-
407 antness were negatively related to unusualness and
408 positively related to prototypicality. Assuming preference
409 and pleasantness are positively related, and familiarity and
410 unusualness are negatively related, our results stand in
411 contrast to the pattern for these participants reported. The
412 same experiment, however, was run using highly trained
413 participants also. In contrast to the musically untrained
414 participants, and in accordance with our data, these partic-
415 ipants rated tone sequences as more pleasant when they
416 had a lower degree of prototypicality and were rated as
417 more unusual. It could be argued then that participants in
418 our experiment gained proficiency during exposure with
419 the novel musical system such that their pattern was similar
420 to the pattern found using Western music and highly
421 trained students.

422 Schellenberg et al. (2008) found that the liking of music
423 follows an inverted U-shaped function of exposure if lis-
424 teners focus on the music, whereas liking increases linearly
425 as a function of exposure if participants listen to the music
426 incidentally. At first glance, this result might not be com-
427 patible with our proposition. However, the incidental
428 condition in the study by Schellenberg et al. (2008) ex-
429 posed participants to the music, while they had to press
430 keys and count the number of occurrences of words in a
431 story that they were listening to simultaneously to the
432 music. This led to lower recognition ratings in the inci-
433 dental group. Thus, participants in the incidental group
434 might not have been overexposed to the music yet, leading
435 to the observed linear relationship between exposure and
436 liking. The data from the study by Schellenberg et al.
437 (2008) could be lying on the left part of an inverted
438 U-shaped function of exposure.

439 It should be noted that the stimuli used in the studies
440 described above (Loui et al. 2010; Schellenberg et al. 2008;
441 Szpunar et al. 2004) were different from the stimuli used
442 in our experiment. In a series of three experiments, Szpunar
443 et al. (2004) used short tone sequences of 5–9 piano tones
444 and excerpts of “classical” music pieces. The number of
445 occurrences of those stimuli was varied to manipulate ex-
446 posure. Thus, stimuli were repeated. Schellenberg et al.
447 (2008) used excerpts of “classical” music pieces. Again,
448 the number of occurrences of those excerpts was changed
449 to manipulate exposure. These two studies differ from ours
450 and the studies by Kuhn and Dienes (2005) and Loui et al.
451 (2010) in that they manipulated exposure by repeating
452 stimuli. Furthermore, Schellenberg et al. (2008) and

453 Szpunar et al. (2004) were not interested in the ability of
454 participants to passively learn a new musical system.

455 Loui et al. (2010) generated tone sequences from a novel
456 musical system as we did. However, they used the Bohlen–
457 Pierce scale. The statistical regularities in their system
458 were based on transitional probabilities. The stimuli in the
459 experiments reported by Kuhn and Dienes (2005) used
460 non-local dependencies. However, our study used Grego-
461 rian chant modes defined by first-order probabilities. Both
462 Kuhn and Dienes (2005) and Loui et al. (2010) claim that
463 participants in their studies were able to generalize the
464 rules on which the music was based. When explicitly asked
465 to do so, participants in the study by Kuhn and Dienes
466 (2005) were unable to differentiate between tone sequences
467 that were similar to the tone sequence set to which they had
468 been exposed and tone sequences that violated the rules on
469 which the other tone sequences had been based. However,
470 they differentiated between those two types of tone se-
471 quences when asked how much they liked them. On the
472 other hand, in the experiments by Loui et al. (2010), and
473 similarly in our study, participants with little or no music
474 training (less than five years) were able to generalize the
475 statistical rules when asked to recognize tone sequences or
476 when asked about their familiarity with the tone sequences.
477 But while recognition or familiarity ratings indicate that
478 participants passively learned about a new musical system,
479 Loui et al. (2010) found no preference for previously heard
480 tone sequences or tone sequences composed in the exposed
481 style over tone sequences following another set of transi-
482 tional probabilities. Similarly, we found no preference for
483 tone sequences composed in the exposed mode over tone
484 sequences of a different mode. Furthermore, there was no
485 significant linear contrast of distinctiveness in the prefer-
486 ence group. So, even though the tone sequences were more
487 or less violating potentially formed expectancies, there was
488 no difference in affect. These differences in the observed
489 pattern between the different studies may be explained
490 when considering the time for which participants were
491 exposed to the tone sequences that were supposed to set
492 forth the rules. While our participants and the participants
493 in the study by Loui et al. (2010) were exposed to the novel
494 music system for about 25 min, participants in the study by
495 Kuhn and Dienes (2005) were “trained” on the rules for
496 six min. The shorter exposure could mean that the boredom
497 effect (Cantor 1968) has not set in yet.

498 In conclusion, it can be said that our results support the
499 idea that passive statistical learning of probability profiles
500 could be involved in gaining musical knowledge, as partic-
501 ipants indicated that tone sequences generated from a
502 probability profile at the high level of distinctiveness, i.e. a
503 “caricatured” version of the probability profile to which
504 they were exposed, were most familiar. Participants in our

study did not prefer tone sequences that other participants
found familiar after the same amount of exposure. Future
explorations might manipulate the length of exposure to
determine whether participants did not prefer the tone se-
quences because they have been overexposed.

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